

CENTRAL LIMIT THEOREM

Here is a look at a strong assumptions version of the CLT. The CLT works under weaker assumptions, but its proof is more difficult.

Theorem 1 *If X_1, X_2, \dots are independent and identically distributed like a random variable X with finite mean, μ , and finite variance σ^2 , then*

$$\lim_{n \rightarrow \infty} \mathbb{P} \left(\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq t \right) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^t e^{-u^2/2} du = \Phi(t) \quad (1)$$

In other words, $(S_n - n\mu)/(\sigma\sqrt{n}) \xrightarrow{d} Z$ or $(\bar{X} - \mu)/(\sigma/\sqrt{n}) \xrightarrow{d} Z$ where $Z \sim N(0, 1)$.

Proof 1 *Let*

$$Z_n = \sum_{i=1}^n \frac{X_i - \mu}{\sigma\sqrt{n}} \quad (2)$$

Since $(X_i - \mu)/(\sigma\sqrt{n})$, $i = 1, 2, \dots$ are independent random variables, and since the mgf of the sum of independent random variables is equal to the product of their mgf's, we can write

$$M_{Z_n}(s) = \prod_{i=1}^n \mathbb{E} \left[\exp \left(s \frac{(X_i - \mu)}{\sigma\sqrt{n}} \right) \right] \quad (3)$$

$$= \left(\mathbb{E} \left[e^{s/(\sigma\sqrt{n})(X-\mu)} \right] \right)^n \quad (4)$$

since the random variables are iid. Hence

$$M_{Z_n}(s) = \left[M_{X-\mu} \left(\frac{s}{\sigma\sqrt{n}} \right) \right]^n \quad (5)$$

Now from before [see mgf notes] we know that

$$M_{X-\mu}(s) = 1 + s\mathbb{E}(X - \mu) + \frac{s^2}{2!}\mathbb{E}[(X - \mu)^2] + \dots + \frac{s^r}{r!}\mathbb{E}[(X - \mu)^r] + \dots \quad (6)$$

and since $\mathbb{E}(X - \mu) = 0$ and $\mathbb{E}[(X - \mu)^2] = \sigma^2$ we have

$$M_{X-\mu}(s) = 1 + \sigma^2 \frac{s^2}{2!} + \dots + \mathbb{E}[(X - \mu)^r] \frac{s^r}{r!} + \dots \quad (7)$$

Thus,

$$M_{X-\mu} \left(\frac{s}{\sigma\sqrt{n}} \right) = 1 + \sigma^2 \frac{s^2}{2n\sigma^2} + \theta_n \frac{s^2}{2n\sigma^2} \quad (8)$$

where $\theta_n \rightarrow 0$ as $n \rightarrow \infty$. So

$$M_{Z_n}(s) = \left(1 + \frac{s^2}{2n} + \frac{\theta_n s^2}{2n\sigma^2} \right)^n \quad (9)$$

$$= \left(1 + \frac{\frac{s^2}{2} + \frac{\theta_n}{2} \cdot \frac{s^2}{\sigma^2}}{n} \right)^n \quad (10)$$

Now let $c_n = \frac{\theta_n s^2}{2\sigma^2}$ and recall that

$$\lim_{n \rightarrow \infty} \left(1 + \frac{a + c_n}{n} \right)^n = e^a$$

if $\lim_{n \rightarrow \infty} c_n = 0$. We now have

$$\lim_{n \rightarrow \infty} M_{Z_n}(s) = e^{s^2/2} \quad (11)$$

But, $e^{s^2/2}$ is the mgf of a standard normal random variable.